Environmental tobacco smoke exposure and non-syndromic orofacial cleft: Systematic review and meta-analysis

Heba J. Sabbagh¹, Khlood K. Baghlaf¹, Hattan M. H. Jamalellail^{2,3}, Abdullah S. Bakhuraybah², Salem M. AlGhamdi², Omar A. Alharbi², Khalid M. AlHarbi², Mona H. A. Hassan⁴

ABSTRACT

INTRODUCTION Environmental tobacco smoke (ETS) is associated with several congenital anomalies, including non-syndromic orofacial clefts (NSOFCs). This systematic review aimed to update the literature on the association between ETS and NSOFCs.

METHODS Four databases were searched up to March 2022, and studies that evaluated the association between ETS and NSOFCs were selected. Two authors selected the studies, extracted the data, and evaluated the risk of bias. Comparing the association of maternal exposure to ETS and active parental smoking with NSOFCs allowed for the creation of pooled effect estimates for the included studies.

RESULTS Twenty-six studies were deemed eligible for this review, of which 14 were reported in a previous systematic review. Twenty five were case-control studies, and one was a cohort study. In total, these studies included 2142 NSOFC cases compared to 118129 controls. All meta-analyses showed an association between ETS and the risk of having a child with NSOFC, based on the cleft phenotype, risk of bias, and year of publication, with a pooled increased odds ratio of 1.80 (95% CI: 1.51–2.15). These studies had a marked heterogeneity, which decreased upon subgrouping based on the recent year of publication and the risk of bias. **CONCLUSIONS** ETS exposure was associated with more than a 1.5-fold increase in the risk of having a child with NSOFC, showing a higher odds ratio than paternal and maternal active smoking.

TRIAL REGISTRATION The study is registered on the International Prospective Register of Systematic Reviews database # CRD42021272909.

Tob. Induc. Dis. 2023;21(June):76

https://doi.org/10.18332/tid/163177

INTRODUCTION

Smoking has been a controversial topic for decades; it remains one of the leading causes of lung cancer in men and breast cancer in women^{1,2}. Smoking may be active or passive. According to the World Health Organization, active smoking is defined as smoking at least one cigarette a day. In contrast, passive smoking is the inhalation of tobacco smoke, also known as secondhand smoke or environmental tobacco smoke (ETS)³.

A recent study reported that ETS increased the risk of developing cardiovascular diseases by 28% and its associated mortality rate by 12%. Individuals affected by ETS are exposed to tobacco smoke at home, at work, and in public places⁴. ETS and active smoking have also been positively associated with congenital anomalies such as neural tube defects⁵, congenital heart defects⁶, and non-syndromic

AFFILIATION

 Department of Pediatric
 Dentistry, Faculty of Dentistry, King Abdulaziz University, Jeddah, Saudi Arabia
 Faculty of Dentistry, King Abdulaziz University, Jeddah, Saudi Arabia
 Primary Health Care, Jizan Department, Ministry of National Guard Health Affairs, Jeddah, Saudi Arabia
 Department of Biostatistics, High Institute of Public Health, Alexandria University, Alexandria, Egypt

CORRESPONDENCE TO

Khlood K. Baghlaf. Department of Pediatric Dentistry, Faculty of Dentistry, King Abdulaziz University, P.O. Box 80200, Jeddah 21589, Saudi Arabia. E-mail: kbaghlaf@kau.edu.sa ORCID ID: https://orcid. org/0000-0002-0326-9633

KEYWORDS

cleft lip, cleft palate, orofacial cleft, passive smoking, environmental tobacco smoke

Received: 24 January 2023 Revised: 26 March 2023 Accepted: 11 April 2023 orofacial clefts (NSOFCs)⁷⁻⁹.

Syndromic orofacial clefts (OFC) are associated with structural or developmental defects, whereas NSOFCs are isolated and unrelated to other abnormalities¹⁰. This condition affects the quality of life; many patients with OFC develop depression, anxiety, lack of self-esteem¹¹, speech defects, facial deformities, and several dental problems, including malocclusion¹². The treatment of OFCs necessitates a multidisciplinary approach, with treatment ranging from infancy through late adolescence¹³.

Globally, the prevalence of NSOFCs is 1.25 per 1000 live births^{8,14}. In 2004, a systematic review and metaanalysis, including 24 case-control studies, evaluated the association between maternal active smoking and the risk of having a child with NSOFC. They reported a modest dose-response effect for cleft lip with or without cleft palate¹⁵. In 2014, a systematic review and metaanalysis, including 14 case-control studies, evaluated the association between ETS and the risk of having a child with NSOFC and reported a positive odds ratio⁸. They also recommended further investigation to provide solid grounds for nicotine exposure⁸.

Since then, many studies have assessed the association between ETS and NSOFCs, and there is a need to update newly published evidence and evaluate the current evidence. Therefore, this systematic review and meta-analysis aimed to update the previous systematic review that pooled studies published up to 2013 by evaluating and comparing the evidence that investigates the association between maternal ETS exposure and NSOFCs in recent studies and published meta-analyses. In addition, it evaluates and compares paternal smoking with ETS exposure, which was not previously assessed.

METHODS

The Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines¹⁶ were followed, and the findings were reported according to the PRISMA statement¹⁷.

Information sources and search strategy

All relevant studies from 1980 to 2022 were identified. A comprehensive search of electronic databases, PubMed, Web of Science, Scopus, and ScienceDirect, between 2013 and March 2022. Studies published before 2013 were identified and recruited from the previous systematic review⁸. The search was not limited to studies published in English-language articles. A manual search of reference lists from identified published work and Google Scholar were also used to search for potentially eligible studies. Medical Subject Headings keywords were used to build a comprehensive search query. Repeated studies were detected and deleted using the EndNote reference manager (EndNote[®] version 9, Niles Software, USA).

The following search terms were used: [(cleft lip) OR (cleft palate) OR (orofacial cleft)] AND [(passive smoking) OR (tobacco smoke pollution) OR (environmental tobacco smoke pollution) OR (smoking)].

Two researchers (AB and SG) were involved in the search strategy. All titles were independently reviewed by two researchers (OA and KH). All duplicates were excluded. Case-control, cohort, and cross-sectional studies that investigated the association between ETS and NSOFCs were included. Studies associated with syndromic OFCs, those that measured active smoking only, and those including genetic models were excluded.

Eligibility criteria

Studies included in this review were selected in accordance with the PICO elements¹⁸: **P**articipants (studies assessing the etiology of NSOFCs), Intervention or exposure (ETS), Comparison (healthy children without OFC), and **O**utcomes (NSOFCs).

Study design

The inclusion criteria included case-control, cohort, and cross-sectional studies investigating the association between ETS and NSOFCs. Studies with a design other than the types mentioned in the inclusion criteria, those associated with syndromic OFCs, those that measured active smoking only, and those including genetic models, were excluded.

Other studies, such as editorials, letters to the editor, pilot studies, historical and literature reviews, *in vitro* studies, and descriptive studies, including case reports and case series, were also excluded.

Study selection and data extraction

Two reviewers (KB and HJ) independently assessed the titles and abstracts of all the identified studies to determine if they met the inclusion criteria. The full-text articles of the selected studies were independently assessed by the same reviewers. Any disagreement between the two reviewers was resolved by consulting a third reviewer (HS). For studies performed on the same sample, studies with additional data were chosen. Two reviewers (KB and HS) assessed the selected articles using a standardized protocol, and the extracted data were recorded in a specific extraction datasheet. The extracted data included author names and citations, site, country, duration of data collection, study design, reported period of maternal exposure, total sample size, percentage of non-smoking mothers exposed to passive smoking and total non-smoking mothers, reported adjusted p-value, and adjusted odds ratio (95% confidence intervals [CI]) for passive smoking, maternal smoking and total sample size, and paternal smoking and total sample size.

Quality assessment and the risk of bias

Both cohort and case-control studies in this review were assessed independently using NOS¹⁹. The scale has a minimum score of 0 and a maximum of 9. It measures the selection of the cases, controls, and cohorts and how they represent the general population; the compatibility of cases, controls, and cohorts based on design and analysis; the exposure ascertainment of case controls; and the outcome of cohorts and the adequacy of their follow-up period. Studies that scored >6, 4–6, and ≤ 3 showed a low, moderate, and high risk of bias, respectively. Studies of moderate and high methodological quality (>5 stars) were included in the meta-analysis⁸. In case of any discrepancy between the two authors, the values were discussed until agreement. The level of agreement between the two authors was evaluated using the kappa score. Grading of Recommendations, Assessment, Development and Evaluations (GRADE) was used to summarize and assess the confidence of evidence and the strength of recommendations. It consists of five domains: risk of bias, inconsistency, indirectness, imprecision, and publication bias.

Data synthesis

Data from the included studies were compiled. The data were organized according to author names and citations, site, duration of the data collection, study design, reported period of maternal ETS exposure, total sample size, percentage of mothers exposed to ETS with their respective p-values and odds ratios, percentage of maternal smoking and total sample size with their respective odds ratios, percentage of paternal smoking, and total sample size with their respective odds ratios, and risk of bias. If needed, a meta-analysis of the association between NSOFCs and ETS was performed.

Both quantitative and qualitative syntheses were performed wherever possible. Studies that compared the association of ETS with different cleft subtypes, including cleft lip (CL) and cleft palate (CP), were presented separately. Quantitative synthesis requires a minimum of two investigations. RevMan was used to conduct the meta-analysis (version 5.1; Nordic Cochrane Center, Cochrane Collaboration, 2001). Cochran's test and Higgin's I² index were used to check for study heterogeneity. I² statistic was classified into moderate heterogeneity (30% to 60%), substantial heterogeneity (50% to 90%) and considerable heterogeneity (75% to 100%)²⁰. When there was a study's heterogeneity, a random-effects model was conducted.

Sensitivity analyses based on subgroups were carried out according to the quality of the studies, the year of publication, and cleft types. Additionally, we assessed active parental smoking in the included studies and compared the results. The formal method of combining individual study data was the odds ratio for individual studies. Subgroup differences were tested using chi-squared. A funnel plot was used to visually assess the probability of small-study effects. Egger's test was used to evaluate publication bias. The significance level was set at p<0.05.

Additionally, meta-regression analysis was performed using Meta-DiSc version 1.4 (<u>http://www.hrc.es/investigacion/metadisc_en.htm</u>) to assess the possible effects of the year of publication, the quality of the study, and the type of smoking on the association between ETS and NSOFCs. All variables entered in the model were binary.

RESULTS

Study selection

The search results from the databases yielded 1081 eligible titles. After the removal of duplicate results, only 821 articles remained. After screening the titles and

reviewing the abstracts, only 21 full-text articles were obtained for comparison that met the inclusion and exclusion criteria. Of these, nine studies were excluded because of: a lack of specification regarding whether mothers were exposed to passive or active smoking (seven studies), focusing on the genetic effects of smoking on newborns (one study) and an overlapping population (one study) (Supplementary file Figure 1). Finally, 12 articles met the inclusion criteria for this systematic review and were suitable for inclusion in the qualitative synthesis (Supplementary file Table 1).

Additionally, 14 studies^{7,9,15,21-30} from the primary systematic review⁸ were included to update the review of this topic and study its effects over multiple decades. Totally, 26 studies were deemed eligible for this review; 25 were case-control studies^{7,9,15,21-40}, and one was a prospective cohort study⁴¹. The data from the two centers were presented separately in a systematic review by Pi et al.^{15,22,26,29,35,37,42} (2018 A and B). To collect data on smoking, all of the included studies used self-reported questionnaires. Casecontrol studies were population-based, hospital-ba sed^{7,9,23,24,27,30,31,33,38,39}, or multi-hospital-based^{8,11,32,33} (Table 1).

ETS and NSOFCs

The definition of maternal ETS exposure was similar in all the included studies. However, Hao et al.³⁴ defined it as exposure to smoke of more than one cigarette per day, either at work or at home. No definition was found in the study by Junaid et al.³⁶, whereas Dien et al.³² used a cutoff point of 15 minutes to count as exposure.

Most studies have used the first trimester as the

Figure 1. Forest plot for meta-analysis of the association between the risk of having an infant with NSOFC and maternal environmental tobacco exposure sub-grouped according to year of publication before and after 2013

	NSO		Con			Odds Ratio		Odds Ratio
tudy or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI Y	ear	M-H, Random, 95% Cl
.3.1 <2013								
Beaty et al 2001	24	107	18	130	2.9%	1.80 [0.92, 3.53] 2		
ittle et al 2004	67	154	111	189	3.8%	0.54 [0.35, 0.83] 2	004	
Ionein et al 2007	235	1227	554	2699	4.6%	0.92 [0.77, 1.09] 2	007	4
Chevrier et al 2008	97	173	70	167	3.8%	1.77 [1.15, 2.72] 2		
ie et al 2008	90	334	106	520	4.2%	1.44 [1.04, 1.99] 2	008	
eite and Koifman 2009	166	274	281	548	4.3%	1.46 [1.09, 1.96] 2	009	
Vang et al 2009	168	586	192	1172	4.5%	2.05 [1.62, 2.60] 2	009	
i et al 2010	59	88	348	651	3.6%	1.77 [1.11, 2.84] 2	010	
hang et al 2010	224	323	169	454	4.2%	3.82 [2.82, 5.17] 2	010	-
ianyan et al 2010	121	200	87	200	3.9%	1.99 [1.34, 2.96] 2	010	
i et al 2011	69	162	54	204	3.7%	2.06 [1.33, 3.20] 2	011	
/irilas et al 2011	34	35	25	35	0.6%	13.60 [1.63, 113.25] 2	011	· · · · ·
ia et al 2011	402	713	27	221	3.8%	9.29 [6.05, 14.26] 2	011	
aghavi et al 2012 Subtotal (95% CI)	113	300 4676	80	300 7490	4.1% 52.0%	1.66 [1.18, 2.35] 2 1.92 [1.35, 2.71]	012	 ◆
otal events	1869		2122					
leterogeneity: Tau ² = 0.38	B: $Chi^2 = 1$	73.92. d	f = 13 (P	< 0.0000	1): $l^2 = 939$	6		
est for overall effect: Z =	,							
.3.2 >2013								
lao et al 2015	285	499	175	480	4.4%	2.32 [1.80, 3.00] 2	015	-
abbagh et al 2015	45	204	47	244	3.7%	1.19 [0.75, 1.88] 2	015	+
loyt et al 2016	148	1102	369	3324	4.6%	1.24 [1.01, 1.52] 2	016	· ·
Cummet et al 2016	1312	4508	2310	9626	4.8%	1.30 [1.20, 1.41] 2	016	•
Ackinney et al 2016	41	95	20	95	3.0%	2.85 [1.50, 5.39] 2	016	
Dien et al 2017	67	170	43	170	3.7%	1.92 [1.21, 3.05] 2	017	
Goveas et al 2017	74	125	53	125	3.5%	1.97 [1.19, 3.26] 2	017	
unaid et al 2017	24	50	12	50	2.3%	2.92 [1.24, 6.87] 2	017	
Pi et al 2018	140	240	664	1420	4.3%	1.59 [1.21, 2.10] 2	018	-
Pi et al 2018 b	56	101	173	561	3.8%	2.79 [1.81, 4.30] 2	018	
ltoe et al 2019	32	150	38	300	3.4%	1.87 [1.11, 3.14] 2		
Chowchuen et al 2020	14	35	24	70	2.3%	1.28 [0.55, 2.95] 2	020	- -
Sato et al 2021	98	187	46566	94174	4.3%	1.13 [0.84, 1.50] 2		+-
Subtotal (95% CI)	- •	7466		110639	48.0%	1.67 [1.40, 1.99]	C. 1999	◆
otal events	2336		50494					
leterogeneity: Tau ² = 0.06	6; Chi² = 4	5.29, df	= 12 (P <	< 0.00001); l² = 74%			
est for overall effect: Z =	5.74 (P <	0.00001)		5			
otal (95% CI)		12142		118129	100.0%	1.80 [1.51, 2.15]		•
otal events	4205		52616					
leterogeneity: Tau ² = 0.17	7; Chi ² = 2	24.93, d	f = 26 (P	< 0.0000	1); l ² = 88%	6	0.01	0,1 1 10 1
		0.00001					0.01	0.1 1 10 1

Table 1. Characteristics of studies

Authors Year Country	Study design	Cleft type	ETS mothers* n/N (%)	AOR (95% CI)	Maternal smoking n/N (%)	AOR (95% CI)	Paternal smoking n/N (%)	AOR (95% CI)	Confounding variables
Beaty et al. ²¹	Case-control	NSOFC	24/107 (22.4)		27/171 (15.8)		-	-	Smoking, alcohol use, daily
2001 US	2001 US	CL/P	14/73 (19.20)	1.04 (0.067–1.62)	17/91 (19.0)	1.36 (0.68–2.72)			vitamin use, urinary tract infection
05		СР	10/34 (29.4)	1.17 (0.68–2.02)	10/44 (23.0)	1.74 (0.75–4.02)			meeton
		Control	18/130 (13.8)		25/182 (13.7)				
Little et al. ¹⁵	Case-control	NSOFC	67/154 (43.5)		80/190 (42.1)				Maternal smoking
2004 UK	Population- based	CL/P	40/76 (52.6)	0.9 (0.5–1.7)	45/112 (40.0)	1.9 (1.1–3.1)	25/67 (37.0)	1.4 (0.7–2.9)	
UK	UdSCU	СР	27/78 (34.6)	1.1 (0.5–2.2)	35/78 (44.9)	2.3 (1.3–4.1)	11/24 (45.0)	2.2 (0.8–5.9)	
		Control	111/189 (58.7)		59/189 (31)		28/119 (23.5)		
Honein et al.22	Case-control	NSOFC	235/1227 (19.1)				-	-	Folic acid exposure, alcohol
2007 US	Population- based	CL/P	147/699 (21.0)	1.0 (0.8–1.3)	200/1461 (13.6)	1.4 (1.1–1.7)	-	-	use, maternal smoking
		СР	88/528 (22.0)	1.1 (0.8–1.4)	92/1461 (6.2)	1.2 (0.9–1.6)			
		Control	554/2699 (20.5)		679/3390 (20.0)				
Chevrier et	Case-control	NSOFC	97/173 (56.1)	1.8 (1.2–3.4)	27/171 (15.8)	1.1 (0.7–1.9)	-	-	Maternal dietary
al. ⁴⁰ 2008		CL/P	65/119 (54.6)			1.0 (0.5–2.0)			folate intake, alcohol consumption, maternal
France		СР	32/54 (59.3)			1.0 (0.3–3.3)			smoking
		Control	70/167 (41.9)	-	25/182 (13.7)	-	-	-	<u> </u>
Lie et al.27	Case-control	NSOFC	90/334 (26.9)	1.05 (0.55–2.00)	239/432 (55.0)	0.81 (0.45–1.44)	-	-	Cigarette smoking, folic
2008 Norway	Hospital- based	CL/P	58/210 (27.6)			1.82 (0.98–3.39)			acid supplement, dietary folate, multivitamins,
Norway	oascu	СР	32/124 (25.8)			0.29 (0.04–2.26)			alcohol use
		Control	106/520 (20.4)	-	243/763 (31.8)	-	-	-	
Leite and	Case-control	NSOFC	166/274 (60.6)	1.48 (1.09–2.01)	51/274 (18.60)	1.28 (0.87–1.97)	59/274 (21.6)	1.02 (0.75–1.52)	Maternal smoking, alcohol
Koifman ²⁴	Hospital-	CL/P		1.39 (1.01–1.98)		1.59 (1.04–2.44)		1.17 (0.80–1.75)	use
2009 Brazil	based	СР		1.67 (0.90–3.11)		0.82 (0.34–1.79)		0.58 (0.19–1.27)	
		Control	281/548 (51)	-	88/548 (16.10)	1.43 (1.25–1.64)	118/548 (21.4)	-	

Continued

Table 1. Continued

Authors Year Country	Study design	Cleft type	ETS mothers* n/N (%)	AOR (95% CI)	Maternal smoking n/N (%)	AOR (95% CI)	Paternal smoking n/N (%)	AOR (95% CI)	Confounding variables
Wang et al. ²⁹ 2009	Case-control Population-	NSOFC	168/586 (28.7)	2.05 (1.47–2.87)	12/344 (2.0)	1.5 (0.52–4.36)	178/334 (30.4)	1.11 (0.82–1.51)	Maternal illness, medication use, maternal smoking,
China	1	Control	192/1172 (16.4)	2.05 (1.47–2.87)	16/1172 (1.3)	1.5 (0.52-4.36)	330/1172 (28.2)	1.11 (0.82–1.51)	toxic exposures, pesticides, alcohol, radiation therapy
Jianyan et al.23	Case-control	CL/P	121/200 (60.5)	1.72 (1.08–2.74)	-	-	105/200 (52.5)	1.04 (0.65–1.67)	Maternal smoking,
2010 China	Hospital- based	Control	87/200 (43.5)	1.72 (1.08–2.74)	-	-	91/200 (45.5)	1.04 (0.65–1.67)	multivitamins
Li et al. ²⁶	Case-control	CL/P	59/88 (67.0)	2.0 (1.2–3.4)	-	-	-		Maternal flu or fever in
2010 China	Population- based	Control	348/651 (54.0)	2.0 (1.2–3.4)	-	-	-	-	early pregnancy, folic acid
Zhang et al.30	Case-control	NSOFC	224/323 (69.3)		14/300 (4.6)				
2010 China	Hospital- based	CL	79/106 (74.5)	3.71 (1.46–9.40)	4/86 (4.7)	7.00 (1.44–34.13)	40/86 (46.5)	14.64 (4.11–52.13)	
China	Uascu	СР	49/77 (36.3)	2.97 (1.32–7.79)	0/77 (0)	<0.01 (<0.01–999.9)	41/77 (53.2)	37.88 (10.5–36.43)	
		CLP	96/140 (68.6)	1.09 (0.41–2.93)	10/140 (7.2)	5.12 (1.30–20.12)	79/140 (56.4)	33.19 (10.5–04.87)	
		Control	169/454 (37.2)	-	6/454 (1.3)		17/454 (3.7)	-	
Jia et al. ⁷	Case-control	NSOFC	402/713 (56.2)		18/713 (2.5)		435/713 (61.0)		Multivitamins supplement,
2011 China	Hospital- based	CL/P	302/537 (56.2)	9.23 (5.96–14.28)	15/537 (2.7)	3.15 (0.71–13.88)	325/537 (60.5)	1.92 (1.40–2.62)	maternal folic acid use, maternal calcium
China	UdSCU	СР	100/176 (56.8)	9.45 (5.73–15.60)	3/176 (1.70)	1.90 (0.31–11.49)	110/176 (62.5)	2.09 (1.40–3.13)	supplement, folic acid,
		Control	27/221 (12.2)	-	2/221 (0.9)	-	98/221 (44.3)	-	alcohol
Li et al. ²⁵	Case-Control	NSOFC	69/162 (42.6)	3.44 (2.24–5.27)	-	-	-	-	Maternal vitamin intake,
2011 China		Control	54/204 (17.4)	-	-	-	-	-	alcohol use
Mirilas et al.28		CL/P	34/35 (45.7)	1.81 (0.69–4.74)	6/35 (17.1)	0.82 (0.24–2.76)	22/35 (62.8)	1.26 (0.48–3.30)	Disease and drugs,
2011 Greece	Hospital- based	Control	25/3531 (31.4)	-	7/35 (20.0)	-	20/35 (57.1)	-	exposure to environmental pollutants, exposure to chemical contaminants
Taghavi et al.9	Case-control	CL/P	113/300 (37.7)	0.613 (0.43-0.87)	7/300 (2.3)	0.516 (0.34–3.93)	-	-	Supplemental vitamin,
2012 Iran	Hospital- based	Control	80/300 (26.7)	-	5/300 (2.0)	-	-	-	folic acid use, radiation exposure, maternal smoking

Continued

Table 1. Continued

Authors Year Country	Study design	Cleft type	ETS mothers* n/N (%)	AOR (95% CI)	Maternal smoking n/N (%)	AOR (95% CI)	Paternal smoking n/N (%)	AOR (95% CI)	Confounding variables	
Hao et al. ³⁴	Case-control	NSOFC	285/499 (57.11)						Medication use,	
2015 China	Multi-	CL/P	214/362 (59.1)	2.52 (1.90-3.33)	26/362 (7.2)	1.25 (0.72–2.17)	218/362 (60.2)	2.17 (1.65–2.87)	maternal smoking, maternal alcohol	
China	China hospital	СР	71/137 (51.8)	1.87 (1.28–2.75)	9/137 (6.6)	1.14 (0.52–2.47)	82/137 (59.9)	2.14 (1.45–.15)		
		Control	175/480 (36.5)		28/480 (5.8)		197/480 (41.0)			
Sabbagh et al.8	Case-control	NSOFC	45/204 (22.0)	1.18 (0.75–1.87)	6/204 (2.9)	0.6 (0.2–1.7)	74/204 (36.3)	1.01 (0.69–1.48)	Maternal medication	
2015 Saudi Arabia	11 Multi- hospital	CL	10/77 (13.0)	0.64 (0.31–1.34)	2/77 (2.7)	0.71 (0.15–3.37)	17/77 (22.1)	0.51 (0.28–0.92)	use and illness, maternal supplements use, maternal	
Sauui Alaula	nospitai	CLP	21/74 (28.4)	1.68 (0.93–3.06)	4/74 (4.1)	0.99 (0.27–3.71)	33/74 (44.6)	1.38 (0.81–2.33)	stress, maternal domestic	
		СР	14/53 (26.4)	1.52 (0.76–3.03)	0/53 (0)		24/53 (45.3)	1.14 (0.78–2.58)	environmental exposure	
		Control	47/244 (19.3)		10/244 (4.1)		90/244 (36.9)			
Hoyt et al.35	Case-control	NSOFC	148/1102 (13.4)	1.25 (1.09–1.04)					Maternal alcohol, pre-	
2016 US	Population- CL based	CL	39/290 (13.4)	1.41 (1.12–1.81)					pregnancy body mass index folic acid exposure,	
05	Uascu	CLP	62/450 (13.8)	1.16 (0.95–1.51)					multivitamins	
		СР	47/362 (12.9)	1.31 (1.26–1.63)	-	-	-	-		
		Control n=3324	369/3324 (11.1)							
Kummet et	Case-control	NSOFC	1914/9482 (21.1)	1.14 (1.02–1.27)	1030/14134 (7.2)	1.2 (1.11–1.46)			Active smoking exposure,	
al. ³⁷ 2016	Population- based	CL		1.14 (0.93–1.39)		1.52 (1.19–1.94)			alcohol use, supplements containing folic acid	
Norway	UdSCU	CLP		1.11 (0.95–1.29)		1.18 (0.97–1.43)			containing rolic actu	
,		СР		1.18 (1.00–139)		1.25 (1.01–1.55)				
		Control n=9626								
Mckinney et	Case-control	CL/P	41/95 (43.2)	6.52 (1.98– 21.44)	93/95 (98.0)	-	-	-	Maternal smoking,	
al. ³⁸ 2016 Thailand	Hospital- based	Control	20/95 (21.1)	6.52 (1.98– 21.44)	92/95 (97.0)	-	-	-	alcohol, folic acid, multivitamins	
Dien et al.32	Case-control	NSOFC	67/170 (39.4)	1.59 (0.50–5.09)	0/340 (0)	-	-	-	Maternal smoking,	
2017 Vietnam	3 Hospital- based	Control	43/170 (25.2)	-	-	-	-	-	caffeine consumption, alcohol	

Continued

Review Paper

Table 1. Continued

Authors Year Country	Study design	Cleft type	ETS mothers* n/N (%)	AOR (95% CI)	Maternal smoking n/N (%)	AOR (95% CI)	Paternal smoking n/N (%)	AOR (95% CI)	Confounding variables	
Goveas et al.33	Case-control	CL/P	74/125 (59.2)	p=0.008, OR=1.97	-	p=0.498	-	-	Alcohol consumption,	
2017 India	multi- hospital based	Control	53/125 (42.4)	-	-	-	-	-	maternal smoking, multivitamins	
Junaid et al. ³⁶	Case-control	NSOFC	24/50 (48.0)	2.46 (0.99–6.08)	1/50 (2.0)	p=1.00	20/50 (40.0)	p=0.41	Paternal alcohol use,	
2018 India	3 hospital based	Control	12/50 (24.0)	2.46 (0.99–6.08)	1/50 (2.0)	-	16/50 (32.0)	-	paternal tobacco use, maternal tobacco exposure	
Pi et al.42	Case-control	NSOFC	140/240 (58.3)						Maternal fever or flu	
2018 (2002–2011)	Population- based	CL/P	131/225 (58.2)	1.6 (1.2–2.2)	-	-	-	-		
China	Uaseu	СР	9/15 (60.0)	p=0.003 1.6 (1.2-2.2)	-	-	-	-		
		Control	664/1420 (46.8)		-	-	-	-		
Pi et al. ⁴² 2018	Case-Control Population-	CL/P	56/101 (55.4)	p=0.002 2.2 (1.4–3.6)					Maternal fever or flu	
(2011-2016) China	based	Control n=561	173/561 (30.8)							
Altoe et al.39	Case-control	NSOFC	32/150 (45.7)	1.98 (1.17–3.34)	13/150 (8.6)	2.04 (0.94-4.43)	-	-	Alcohol consumption,	
2019 Brazil	Hospital- based	Control	38/300 (54.2)	-	15/300 (5.0)		-		use of medication, diseases	
Chowchuen et al. ³¹	Case-control Hospital-	CL/P	14/35 (40.0)	1.77 (0.52–6.04)	1/34 (2.86)	p=0.624	-	-	Alcohol intake, smoking, vitamin use, calcium,	
2021 Thailand	based	Control	24/70 (34.29)	-	1/70 (1.4)	-	-	-	iron and folic acid	
Sato et al.41	Prospective-	NSOFC	98/187 (52.4)		83/187 (44.4)				Psychological distress,	
2021 Japan	cohort	CL/P	82/146 (56.1)	1.49 (0.93–2.39)	68/146 (46.5)	0.82 (0.34–1.99)			maternal alcohol consumption,	
заран		СР	16/41 (39.0)	-	15/41 (36.5)				maternal active smoking,	
		Control	46566/94174 (49.4)	-	38228/94174 (40.5)				BMI, folic acid	

*Non-smoking mothers exposed to environmental tobacco smoke (ETS). AOR: adjusted odds ratio. N: total sample.

measurement period for maternal passive smoking. However, the study by Honein et al.²² used a period of three months before pregnancy until birth: one year of pre-gestation along with the first trimester²⁴, the first 28 weeks of pregnancy²⁹, and the first month before pregnancy through the end of the first trimester^{28,30}. Some studies used multiple periods of measurement, including one-year pre-gestation and the first trimester and three months pre-gestation and the first trimester^{9,38}. Junaid et al.³⁶ did not mention the measurement period in their study, whereas Pi et al.⁴² used the measurement period from the last menstrual period till the second trimester⁴¹. In this systematic review, we combined exposure, pre-gestation, and the first trimester in the meta-analysis.

Study quality and risk of bias

The included studies were assessed by two authors, AB and SA, and the inter-rater agreement for the evaluation of the risk of bias was very good (Kappa score = 89). The Supplementary file Table 1 shows the included 11 studies distributed according to the NOS risk of bias scores. Out of these, only two were found to have a low risk of bias^{37,40}. The remaining nine studies were found to have a moderate to high risk of bias^{31-36,38,39,41,42}. This was mainly due to the absence of comparability and matching between cases and controls in many studies^{31,33,35,36,42}. Furthermore, the studies showed bias in exposure ascertainment, as it was not possible for interviewers to be blinded to the cases or control status. The NOS descriptions and scores for the included 11 studies are presented in Supplementary file Table 1.

Meta-analysis

Meta-analysis was conducted on 27 studies (considering the Pi et al.⁴² study to have two parts, A and B), which were then sub-grouped to assess sensitivity (Supplementary file Figures 2 and 4 to 8). The analysis included 12142 NSOFC cases and 118129 controls. There was a highly significant relationship between ETS and NSOFCs (p<0.01) with an increased odds ratio of having a child with NSOFC (OR=1.80; 95% CI: 1.51–2.15) (Table 2 and Figure 1).

Year of publication

The forest plot for the relationship between ETS and having a child with NSOFC, sub-grouped based

on the year of publication, showed that studies published after 2013 had increased ETS odds ratio of having a child with NSOFC (OR=1.67; 95% CI: 1.40–1.99), similar to earlier studies. Even though the overall heterogeneity between studies was high (I²=88%), it decreased to 74% in studies published after 2013 compared to those published before 2013. Nevertheless, there were no significant differences between the subgroups (p=0.49) (Figure 1).

NSOFC phenotypes

The forest plot for the relationship between ETS and the risk of having a child with CL or CP showed a highly significant correlation between ETS and both CL or CP and CP (p<0.001 and p=0.01, respectively), with increased odds ratio of 1.85 (95% CI: 1.46–2.34) for CL or CP and 1.80 (95% CI: 1.47–2.21) for CP. After removing the two studies, one with a high odds ratio⁷ and one with an extremely high CI²⁸, the odds ratio remained high and significant (OR=1.57; 95% CI: 1.33–1.84 for CL or CP; and OR=1.44; CI: 1.31– 1.75 for CP) (Figure 2).

Only three studies evaluated the CL or CP subphenotypes^{30,35,40}. Only one of these studies had a low risk of bias⁴⁰. Maternal ETS exposure showed a nonsignificant increase in the OR for both CL (OR=1.61; 95% CI: 0.54–4.82) and CP (OR=1.98; 95% CI: 0.98–4.01).

Maternal and paternal active smoking

The forest plot for the meta-analysis of the association between the risk of having a child with NSOFC

Table 2. Results of meta-analysis subgrouping

Subgroup at	nalysis	OR (95% CI)	р	Heterogeneity $I^2 (\%)$
Publication year	<2013	1.92 (1.35–2.71)	0.0002	93
	>2013	1.67 (1.40–1.99)	<0.001	74
NSOFC types	CL/P	1.85 (1.46–2.34)	<0.001	87
	СР	1.72 (1.13–2.63)	0.01	89
Active smoking	Maternal	1.51 (1.23–1.86)	<0.001	59
	Paternal	1.51 (1.11–2.06)	0.008	79
Risk of bias	Low	1.42 (1.17–1.71)	0.0003	81
	High	2.23 (1.65–3.01)	< 0.001	88

Figure 2. Forest plot for meta-analysis of the association between the risk of having an infant with cleft lip with or without cleft palate (CL/P) or cleft palate (CP) and its association with maternal environmental tobacco smoking

Study or Subgroup	Franks	ental	Con		Malakt	Odds Ratio	Venn	Odds Ratio
	Events	Total	Events	rotai	weight	M-H, Random, 95% CI	rear	M-H, Random, 95% Cl
CL/P	0.000			1022		000000000000000000000000000000000000000		
Beaty et al 2001	14	73	18	130	2.5%	1.48 [0.69, 3.18]		
Little et al 2004	40	76	111	189	3.0%	0.78 [0.46, 1.33]	2004	
Honein et al 2007	147	699	554	2699	3.7%	1.03 [0.84, 1.27]		+
Chevrier et al 2008	65	119	70	167	3.2%	1.67 [1.04, 2.68]		
Lie et al 2008	58	210	106	520	3.4%	1.49 [1.03, 2.16]	2008	-
Leite and Koifman 2009	124	208	281	548	3.5%	1.40 [1.01, 1.94]	2009	
Jianyan et al 2010	121	200	87	200	3.3%	1.99 [1.34, 2.96]	2010	
Li et al 2010	59	88	348	651	3.2%	1.77 [1.11, 2.84]	2010	
Zhang et al 2010	175	246	169	454	3.4%	4.16 [2.97, 5.81]	2010	
Jia et al 2011	302	537	27	221	3.2%	9.23 [5.96, 14.30]	2011	
Mirilas et al 2011	34	35	25	35	0.8%	13.60 [1.63, 113.25]	2011	
Taghavi et al 2012	113	300	80	300	3.4%	1.66 [1.18, 2.35]	2012	
Hao et al 2015	214	362	175	480	3.5%	2.52 [1.90, 3.33]		
Sabbagh et al 2015	31	151	47	244	3.1%	1.08 [0.65, 1.80]	2015	
Hoyt et al 2016	101	740	369	3324	3.6%	1.27 [1.00, 1.60]		++-
Mckinney et al 2016	41	95	20	95	2.8%	2.85 [1.50, 5.39]		
Goveas et al 2017	74	125	53	125	3.1%	1.97 [1.19, 3.26]		
Pi et al 2018 b	56	101	173	561	3.2%	2.79 [1.81, 4.30]		
Pi et al 2018	131	225	664	1420	3.5%	1.59 [1.19, 2.11]		+
Chowchuen et al 2020	14	35	24	70	2.3%	1.28 [0.55, 2.95]		
Sato et al 2021	82	146	46566	94174	3.5%	1.31 [0.94, 1.82]		
Subtotal (95% CI)	01	4771	40500	106607	65.2%	1.85 [1.46, 2.34]	2021	•
Total events	1996		49967					~
			= 20 (P ·	< 0.00001); I* = 87%	,		
Heterogeneity: Tau ² = 0.2 Test for overall effect: Z = CP			= 20 (P 4	< 0.00001); I⁼= 87%			
Test for overall effect: Z = CP	5.13 (P < 0	.00001)					2001	
Test for overall effect: Z = CP Beaty et al 2001	5.13 (P < 0 10	.00001) 34	18	133	2.2%	2.66 [1.09, 6.48]		
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004	5.13 (P < 0 10 27	.00001) 34 42	18 111	133 189	2.2% 2.6%	2.66 (1.09, 6.48) 1.26 (0.63, 2.53)	2004	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007	5.13 (P < 0 10 27 88	.00001) 34 42 528	18 111 554	133 189 2699	2.2% 2.6% 3.6%	2.66 (1.09, 6.48) 1.26 (0.63, 2.53) 0.77 (0.60, 0.99)	2004 2007	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008	5.13 (P < 0 10 27 88 32	.00001) 34 42 528 54	18 111 554 70	133 189 2699 167	2.2% 2.6% 3.6% 2.8%	2.66 (1.09, 6.48) 1.26 (0.63, 2.53) 0.77 (0.60, 0.99) 2.02 (1.08, 3.76)	2004 2007 2008	-
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008	5.13 (P < 0 10 27 88 32 32 32	.00001) 34 42 528 54 124	18 111 554 70 106	133 189 2699 167 520	2.2% 2.6% 3.6% 2.8% 3.2%	2.66 (1.09, 6.48) 1.26 (0.63, 2.53) 0.77 (0.60, 0.99) 2.02 (1.08, 3.76) 1.36 (0.86, 2.14)	2004 2007 2008 2008	-
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010	5.13 (P < 0 10 27 88 32 32 49	.00001) 34 42 528 54 124 77	18 111 554 70 106 169	133 189 2699 167 520 454	2.2% 2.6% 3.6% 2.8% 3.2% 3.1%	2.66 (1.09, 6.48) 1.26 (0.63, 2.53) 0.77 (0.60, 0.99) 2.02 (1.08, 3.76) 1.36 (0.86, 2.14) 2.95 (1.79, 4.87)	2004 2007 2008 2008 2010	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011	5.13 (P < 0 10 27 88 32 32 49 100	.00001) 34 42 528 54 124 77 176	18 111 554 70 106 169 27	133 189 2699 167 520 454 221	2.2% 2.6% 3.6% 2.8% 3.2% 3.1% 3.1%	2.66 [1.09, 6.48] 1.26 [0.63, 2.53] 0.77 [0.60, 0.99] 2.02 [1.08, 3.76] 1.36 [0.86, 2.14] 2.95 [1.79, 4.87] 9.45 [5.73, 15.60]	2004 2007 2008 2008 2010 2011	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015	5.13 (P < 0 10 27 88 32 32 49 100 71	.00001) 34 42 528 54 124 77 176 137	18 111 554 70 106 169 27 175	133 189 2699 167 520 454 221 480	2.2% 2.6% 3.6% 2.8% 3.2% 3.1% 3.1% 3.4%	2.66 [1.09, 6.48] 1.26 [0.63, 2.53] 0.77 [0.60, 0.99] 2.02 [1.08, 3.76] 1.36 [0.86, 2.14] 2.95 [1.79, 4.87] 9.45 [5.73, 15.60] 1.87 [1.28, 2.75]	2004 2007 2008 2008 2010 2011 2015	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015 Sabbagh et al 2015	5.13 (P < 0 10 27 88 32 32 49 100 71 14	.00001) 34 42 528 54 124 77 176 137 53	18 111 554 70 106 169 27 175 47	133 189 2699 167 520 454 221 480 244	2.2% 2.6% 3.6% 3.2% 3.1% 3.1% 3.4% 2.7%	2.66 [1.09, 6.48] 1.26 [0.63, 2.53] 0.77 [0.60, 0.99] 2.02 [1.08, 3.76] 1.36 [0.86, 2.14] 2.95 [1.79, 4.87] 9.45 [5.73, 15.60] 1.87 [1.28, 2.75] 1.50 [0.76, 3.00]	2004 2007 2008 2008 2010 2011 2015 2015	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015 Sabbagh et al 2015 Hoyt et al 2016	5.13 (P < 0 10 27 88 32 32 49 100 71 14 47	.00001) 34 42 528 54 124 77 176 137 53 362	18 111 554 70 106 169 27 175 47 369	133 189 2699 167 520 454 221 480 244 3324	2.2% 2.6% 3.6% 2.8% 3.2% 3.1% 3.1% 3.4% 2.7% 3.5%	2.66 [1.09, 6.48] 1.26 [0.63, 2.53] 0.77 [0.60, 0.99] 2.02 [1.08, 3.76] 1.36 [0.86, 2.14] 2.95 [1.79, 4.87] 9.45 [5.73, 15.60] 1.87 [1.28, 2.75] 1.50 [0.76, 3.00] 1.19 [0.86, 1.65]	2004 2007 2008 2008 2010 2011 2015 2015 2015	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015 Sabbagh et al 2015 Hoyt et al 2016 Pi et al 2018	5.13 (P < 0 10 27 88 32 32 49 100 71 14 47 9	.00001) 34 42 528 54 124 77 176 137 53 362 15	18 111 554 70 106 169 27 175 47 369 664	133 189 2699 167 520 454 221 480 244 3324 1420	2.2% 2.6% 3.6% 2.8% 3.2% 3.1% 3.1% 3.4% 2.7% 3.5% 1.9%	2.66 [1.09, 6.48] 1.26 [0.63, 2.53] 0.77 [0.60, 0.99] 2.02 [1.08, 3.76] 1.36 [0.86, 2.14] 2.95 [1.79, 4.87] 9.45 [5.73, 15.60] 1.87 [1.28, 2.75] 1.50 [0.76, 3.00] 1.19 [0.86, 1.65] 1.71 [0.60, 4.82]	2004 2007 2008 2010 2011 2015 2015 2016 2018	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015 Sabbagh et al 2015 Hoyt et al 2016 Pi et al 2018 Sato et al 2021 Subtotal (95% CI)	5,13 (P < 0 10 27 88 32 32 49 100 71 14 14 47 9 16	.00001) 34 42 528 54 124 77 176 137 53 362	18 111 554 70 106 169 27 175 47 369 664	133 189 2699 167 520 454 221 480 244 3324	2.2% 2.6% 3.6% 2.8% 3.2% 3.1% 3.1% 3.4% 2.7% 3.5%	2.66 [1.09, 6.48] 1.26 [0.63, 2.53] 0.77 [0.60, 0.99] 2.02 [1.08, 3.76] 1.36 [0.86, 2.14] 2.95 [1.79, 4.87] 9.45 [5.73, 15.60] 1.87 [1.28, 2.75] 1.50 [0.76, 3.00] 1.19 [0.86, 1.65]	2004 2007 2008 2010 2011 2015 2015 2016 2018	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015 Sabbagh et al 2015 Hoyt et al 2016 Pi et al 2018 Sato et al 2021 Subtotal (95% CI) Total events	5.13 (P < 0 10 27 88 32 32 49 100 71 14 47 9 16 495	.00001) 34 42 528 54 124 77 176 137 362 15 41 1643	18 111 554 700 106 169 27 175 47 369 664 46566 48876	133 189 2699 167 520 454 221 480 244 3324 1420 94174 104025	2.2% 2.6% 3.6% 2.8% 3.1% 3.1% 3.4% 2.7% 3.5% 1.9% 2.8% 34.8%	2.66 (1.09, 6.48) 1.26 (0.63, 2.53) 0.77 (0.60, 0.99) 2.02 (1.08, 3.76) 1.36 (0.86, 2.14) 2.95 (1.79, 4.87) 9.45 (5.73, 15.60) 1.87 (1.28, 2.75) 1.50 (0.76, 3.00) 1.19 (0.86, 1.65) 1.71 (0.60, 4.82) 0.65 (0.35, 1.23)	2004 2007 2008 2010 2011 2015 2015 2016 2018	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015 Sabbagh et al 2015 Hoyt et al 2016 Pi et al 2018 Sato et al 2021 Subtotal (95% CI) Total events Heterogeneity: Tau ² = 0.4	5,13 (P < 0 10 27 88 32 32 49 100 71 14 47 9 16 495 7; Chi ^e = 98	.00001) 34 42 528 54 124 77 176 137 53 362 15 362 11 1643 3.68, df=	18 111 554 700 106 169 27 175 47 369 664 46566 48876	133 189 2699 167 520 454 221 480 244 3324 1420 94174 104025	2.2% 2.6% 3.6% 2.8% 3.1% 3.1% 3.4% 2.7% 3.5% 1.9% 2.8% 34.8%	2.66 (1.09, 6.48) 1.26 (0.63, 2.53) 0.77 (0.60, 0.99) 2.02 (1.08, 3.76) 1.36 (0.86, 2.14) 2.95 (1.79, 4.87) 9.45 (5.73, 15.60) 1.87 (1.28, 2.75) 1.50 (0.76, 3.00) 1.19 (0.86, 1.65) 1.71 (0.60, 4.82) 0.65 (0.35, 1.23)	2004 2007 2008 2010 2011 2015 2015 2016 2018	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015 Sabbagh et al 2015 Hoyt et al 2016 Pi et al 2018 Sato et al 2021 Subtotal (95% CI) Total events Heterogeneity: Tau ² = 0.4 Test for overall effect. Z = Total (95% CI)	5.13 (P < 0 10 27 88 32 32 49 100 71 14 47 9 16 495 7; Chiᢪ = 9€ 2.51 (P = 0	.00001) 34 42 528 54 124 77 176 137 53 362 15 362 11 1643 3.68, df=	18 111 554 70 106 169 27 175 47 369 664 46566 48876 11 (P <	133 189 2699 167 520 454 221 480 244 3324 1420 94174 104025 0.00001);	2.2% 2.6% 3.6% 2.8% 3.1% 3.1% 3.4% 2.7% 3.5% 1.9% 2.8% 34.8%	2.66 (1.09, 6.48) 1.26 (0.63, 2.53) 0.77 (0.60, 0.99) 2.02 (1.08, 3.76) 1.36 (0.86, 2.14) 2.95 (1.79, 4.87) 9.45 (5.73, 15.60) 1.87 (1.28, 2.75) 1.50 (0.76, 3.00) 1.19 (0.86, 1.65) 1.71 (0.60, 4.82) 0.65 (0.35, 1.23)	2004 2007 2008 2010 2011 2015 2015 2016 2018	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015 Sabbagh et al 2015 Hoyt et al 2016 Pri et al 2018 Sato et al 2021 Subtotal (95% CI) Total events Heterogeneity: Tau ² = 0.4 Test for overall effect: Z = Total (95% CI) Total events	5,13 (P < 0 10 27 88 32 32 49 100 71 14 47 9 16 495 7; Chi ^p = 9£ 2,51 (P = 0 2,491	.00001) 34 42 528 54 124 77 176 137 53 362 15 411 1643 3.68, df= .01) 6414	18 111 554 70 106 169 277 175 47 369 664 46566 48876 11 (P < 98843	133 189 2699 167 520 454 221 480 244 3324 1420 94174 104025 0.00001); 210632	2.2% 2.6% 3.6% 2.8% 3.1% 3.1% 3.1% 3.4% 2.7% 3.5% 1.9% 2.8% 34.8% 1*= 89%	2.66 [1.09, 6.48] 1.26 [0.63, 2.53] 0.77 [0.60, 0.99] 2.02 [1.08, 3.76] 1.36 [0.86, 2.14] 2.95 [1.79, 4.87] 9.45 [5.73, 15.60] 1.87 [1.28, 2.75] 1.50 [0.76, 3.00] 1.19 [0.86, 1.65] 1.71 [0.60, 4.82] 0.65 [0.35, 1.23] 1.72 [1.13, 2.63]	2004 2007 2008 2010 2011 2015 2015 2016 2018	
Test for overall effect: Z = CP Beaty et al 2001 Little et al 2004 Honein et al 2007 Chevrier et al 2008 Lie et al 2008 Zhang et al 2010 Jia et al 2011 Hao et al 2015 Sabbagh et al 2015 Hoyt et al 2016 Pi et al 2018 Sato et al 2021 Subtotal (95% CI) Total events Heterogeneity: Tau ² = 0.4 Test for overall effect. Z = Total (95% CI)	5.13 (P < 0 10 27 88 32 32 49 100 71 14 47 9 16 495 7; Chi [≠] = 9£ 2.51 (P = 0 2491 9; Chi [≠] = 25	.00001) 34 42 528 54 124 77 176 137 53 362 15 41 1643 3.68, df= .01) 6414 3.59, df	18 111 554 70 106 169 277 175 47 369 664 46566 48876 11 (P < 98843	133 189 2699 167 520 454 221 480 244 3324 1420 94174 104025 0.00001); 210632	2.2% 2.6% 3.6% 2.8% 3.1% 3.1% 3.1% 3.4% 2.7% 3.5% 1.9% 2.8% 34.8% 1*= 89%	2.66 [1.09, 6.48] 1.26 [0.63, 2.53] 0.77 [0.60, 0.99] 2.02 [1.08, 3.76] 1.36 [0.86, 2.14] 2.95 [1.79, 4.87] 9.45 [5.73, 15.60] 1.87 [1.28, 2.75] 1.50 [0.76, 3.00] 1.19 [0.86, 1.65] 1.71 [0.60, 4.82] 0.65 [0.35, 1.23] 1.72 [1.13, 2.63]	2004 2007 2008 2010 2011 2015 2015 2016 2018	

and parental active smoking among studies that evaluated ETS showed a significant association with an increased OR between maternal (OR=1.53; 95% CI: 1.23–1.88; p<0.001) and paternal active smoking (OR=1.51; 95% CI: 1.11–2.06; p<0.001) and NSOFCs. Among the studies that evaluated the paternal active smoking effect on NSOFCs, maternal ETS showed a higher OR (OR=2.21; 95% CI: 1.42–3.45). However, there was no significant difference between maternal ETS and paternal smoking subgrouping (p=0.17) (Supplementary file Figures 3 and 4).

Sensitivity test

To demonstrate the stability and reliability of the meta-analysis results, a sensitivity analysis was conducted between different study subgroups according to cleft phenotype (Figure 2), the exclusion of studies with extreme results (Supplementary file Figure 6), the risk of bias (Supplementary file Figure 5) and the period of maternal ETS exposure (Supplementary file Figure 8). All meta-analyses showed consistent results of a significant association between maternal ETS exposure and an increased OR

Table 3. GRADE profile

			Numl patie		Effect	Quality				
Number of studies	Design	Risk of bias	Inconsistency	Indirectness	Imprecision	Other considerations	NSOFC*	Control		
26	Case- control and cohort ^a	Serious ^b	Not serious ^c	Not serious ^d	Not serious ^e	Confounding ^f	11798	117985 ⁹	Pooled ^h	⊕⊕⊕⊖ Moderate

Question: Is the environmental tobacco smoking associated with the risk of having an infant with NSOFC in all included twenty-six studies? Setting: general infant population. *NSOFC (case-control studies); measured with standard indices; better indicated by lower values. a Twenty-six case-control studies and one cohort. Sato et al. 2021 investigated the association between maternal ETS and the risk of having an infant with NSOFC in all included twenty-six studies. The OR ranged from 0.54 to 9.29. b In Hoyt et al. 2016, Mckinney et al. 2016, Dien et al. 2017, Goveas et al. 2017, Junaid et al. 2018, and Pi et al. 2018 there was a serious risk of bias; therefore, there was a downgrading for risk of bias. c There was no evidence of inconsistency. Most studies showed a significant association between NSIFC and ETS. Therefore, no downgrading was done for this inconsistency. d Data were not downgraded for indirectness because all case-control studies were conducted worldwide. e No downgrading for imprecision because all confidence intervals were narrow and no overlaps. f No downgrading due to the plausible confounding was done; most studies controlled for the other confounding factors such as patient cooperation, isolation of the tooth and type of the teeth (upper or lower molars). g Total number of infants from the 27 studies. h Pooling of meta-analysis z=6.23, p=0.000001 with high heterogeneity.

for the risk of having a child with NSOFC. However, none of the sub-grouped studies accounted for substantial heterogeneity between the studies. All meta-analyses showed significant heterogeneity of $\geq 75\%$.

Evaluation of small-study effects

The funnel plots for all included studies that assessed the relationship between NSOFC phenotypes (CL or CP and CP) and passive smoking did not have the shape of a funnel. However, it was almost a symmetrical funnel plot around the central line, which indicates a publication bias (Supplementary file Figures 9 and 10). However, Egger's test detected no publication bias.

Stability of the evidence

The cumulative meta-analysis figure shows the stability of the evidence from 2011 to 2021. Regarding sufficiency ('Are additional studies needed to establish the existence of the phenomenon?'), from the beginning of this cumulative meta-analysis, the 95% CIs around the OR included the final average effect size (ES) obtained at the end of the cumulative meta-analysis. Regarding stability ('Will additional studies change the aggregate picture of the phenomenon?'), from the beginning of this cumulative meta-analysis, the mean ES appeared to be stable. Therefore, it would be difficult to argue that a subsequent study would alter the emergent picture of this effect beyond

the evidence that the first few studies have produced (Supplementary file Figure 10).

Meta-regression random effects model

A significant model implies that ES is associated with the variables (Table 3). There is a significant difference within groups, which shows that there may be more variables associated with ES. ES was lower for studies conducted after 2013 than for those before 2013. ES was greater in low-quality studies than in high-quality studies. Smoking, whether active or passive, did not affect ES after controlling for other variables in the meta-analysis model. R^2 indicated that 33% of the heterogeneity was accounted for by the addition of predictors to the model compared to an 'empty' model. In other words, this represented the percentage of heterogeneity explained by grouplevel variables in the model (Supplementary file Figure 11).

DISCUSSION

A systematic review conducted in 2015 assessed the association between maternal ETS exposure and NSOFCs that included studies published between 1980 and 2013⁸. However, paternal smoking in ETS studies and the association between ETS and CL/P subphenotypes (CL vs CLP) were not sufficiently discussed. Therefore, this systematic review was conducted to update the literature. This study consistently suggests a more than 1.5 increase in the

risk of NSOFC phenotypes associated with paternal active smoking. A corresponding value of 2.21 was observed with ETS exposure.

In 2020, the World Health Organization highlighted a wide range of adverse health effects of nicotine exposure on infant and child development that result from ETS. They have urged protective policies directed toward smoke-free generations⁴³. Maternal ETS exposure is associated with multiple birth defects and stillbirth⁴⁴. ETS has been reported to cause fetal hypoxia, which leads to fetal growth retardation^{45,46}. In this systematic review, the association and OR between maternal ETS exposure and having a child with NSOFC were significant. The difference between the outcomes of the studies published before and after 2013 was not significant. However, there was a small decrease in the OR (from 1.92 to 1.67), a smaller 95% CI range [from (1.35-2.71) to (1.40-1.99)], and heterogeneity ($I^2 = 93\%$ to 74%), which could be related to the recent improvement in study design and data collection method. Additionally, the studies published after 2013 showed less heterogeneity.

Furthermore, paternal smoking could be associated with having a child with NSOFC either directly by affecting sperm development or indirectly by increasing maternal ETS exposure⁴⁷⁻⁴⁹. Our findings suggest a possible association between ETS and NSOFCs (OR=2.21) that is stronger than that with active paternal smoking (OR=1.51) which could support the indirect effect of paternal smoking by increasing maternal ETS exposure. However, as the difference between the two associations was not significant, this suggestion needs further investigation to verify it.

In this systematic review, the OR of maternal ETS exposure and having a child with CL or CP was higher than that of having a child with CP (Figure 2). However, in the old systematic review, the OR of ETS and having a child with CP was similar to that of having a child with CL or CP⁸. Additionally, this study further assessed the NSOFC phenotype by including three articles investigating ETS association with CL, CLP, and CP formation. The study found an association between an increased risk of CL and CLP, though the available information was inadequate for reporting significant findings. Thus, further studies are needed to evaluate the effects of ETS on different NSOFC subphenotypes.

Our study supports the importance of implementing smoke-free legislation. In England and Northern Ireland, a study assessed the impact of smoke-free legislation on the prevalence of NSOFC. A reduction of 37% and 8%, respectively, in smoking was detected among active female smokers between 2000 and 2018⁵⁰. Although they found no significant reduction in NSOFC prevalence, their results highlight the importance of public health measures, including smoke-free legislation in restaurants and prevention programs among pregnant females in controlling active smoking⁵¹.

The current worldwide response to the coronavirus disease 2019 (COVID-19) was a significant and widespread effect on stress and psychological conditions⁵². The COVID-19 pandemic has influenced the lifestyles of individuals, affecting their nicotine use and exposure^{53,54}. However, a slight decrease in ETS was reported due to the lockdown⁵³. This systematic review did not find any studies that evaluated nicotine exposure in NSOFCs after the COVID-19 pandemic. Therefore, future studies should evaluate this period.

Limitations

Even though our meta-analysis included 26 articles, there were a few limitations. The heterogeneity between studies and the restricted high-quality case-control studies were the two main limitations. Thus, meta-analyses of some findings might lack adequate power and not allow accurate evaluation of heterogeneity with small-study effects and reporting biases. Combining evidence is also more challenging in the presence of different confounding variables, such as the frequency and distance of passive smoking. Moreover, there is still limited evidence supporting the effect of ETS on the development of different NSOFC sub-phenotypes and severity. These limitations and gaps in the literature highlight the need for well-conducted cohort studies that consider the definition of passive smoking and the evaluation of nicotine exposure using a validated, exact method instead of a subjective method like a questionnaire.

CONCLUSIONS

There was a highly significant association between maternal ETS exposure and NSOFCs in children, indicating the importance of implementing smoke-free legislation and maternal pregnancy care. However, the included studies showed marked heterogeneity. Future case-control studies to examine the association between ETS exposure and NSOFCs should consider the definition of ETS and the evaluation of nicotine exposure using an objective measuring tool.

REFERENCES

- Macacu A, Autier P, Boniol M, Boyle P. Active and passive smoking and risk of breast cancer: a meta-analysis. Breast Cancer Res Treat. 2015;154(2):213-224. doi:<u>10.1007/</u> <u>s10549-015-3628-4</u>
- O'Keeffe LM, Taylor G, Huxley RR, Mitchell P, Woodward M, Peters SAE. Smoking as a risk factor for lung cancer in women and men: a systematic review and metaanalysis. BMJ Open. 2018;8(10):e021611. doi:<u>10.1136/ bmjopen-2018-021611</u>
- 3. International Agency for Research on Cancer Working Group on the Evaluation of Carcinogenic Risks to Humans. Tobacco smoke and involuntary smoking. IARC Monogr Eval Carcinog Risks Hum. 2004;83:1-1438.
- Khoramdad M, Vahedian-Azimi A, Karimi L, Rahimi-Bashar F, Amini H, Sahebkar A. Association between passive smoking and cardiovascular disease: a systematic review and meta-analysis. IUBMB Life. 2020;72(4):677-686. doi:<u>10.1002/iub.2207</u>
- Meng X, Sun Y, Duan W, Jia C. Meta-analysis of the association of maternal smoking and passive smoking during pregnancy with neural tube defects. Int J Gynaecol Obstet. 2018;140(1):18-25. doi:10.1002/ijgo.12334
- Mamun MAA, Hussain M, Khan MKES. Passive smoking and congenital heart defects in offspring among Bangladesh. Malays J Paediatr Child Health. 2021;27(1):56-61. doi:<u>10.51407/mjpch.v27i1.120</u>
- Jia ZL, Shi B, Chen CH, Shi JY, Wu J, Xu X. Maternal malnutrition, environmental exposure during pregnancy and the risk of non-syndromic orofacial clefts. Oral Dis. 2011;17(6):584-589. doi:10.1111/j.1601-0825.2011.01810.x
- Sabbagh HJ, Hassan MH, Innes NP, Elkodary HM, Little J, Mossey PA. Passive smoking in the etiology of nonsyndromic orofacial clefts: a systematic review and metaanalysis. PLoS One. 2015;10(3):e0116963. doi:<u>10.1371/</u> journal.pone.0116963
- Taghavi N, Mollaian M, Alizadeh P, Moshref M, Modabernia Sh, Akbarzadeh AR. Orofacial clefts and risk factors in tehran, Iran: a case control study. Iran Red Crescent Med J. 2012;14(1):25-30.
- 10. McGarry A. The influence of genetics on syndromic and non-syndromic cases of cleft lip and cleft palate. Accessed April 11, 2023. <u>https://publichealth.gwu.edu/sites/ default/files/images/Alice%20McGarry_PDF.pdf</u>
- Kawalec A, Nelke K, Pawlas K, Gerber H. Risk factors involved in orofacial cleft predisposition - review. Open Med (Wars). 2015;10(1):163-175. doi:10.1515/med-

2015-0027

- Rando GM, Jorge PK, Vitor LLR, et al. Oral healthrelated quality of life of children with oral clefts and their families. J Appl Oral Sci. 2018;26:e20170106. doi:10.1590/1678-7757-2017-0106
- Sischo L, Wilson-Genderson M, Broder HL. Qualityof-Life in children with orofacial clefts and caregiver well-being. J Dent Res. 2017;96(13):1474-1481. doi:10.1177/0022034517725707
- Mossey PA, Modell B. Epidemiology of oral clefts 2012: an international perspective. Front Oral Biol. 2012;16:1-18. doi:10.1159/000337464
- Little J, Cardy A, Arslan MT, et al. Smoking and orofacial clefts: a United Kingdom-based case-control study. Cleft Palate Craniofac J. 2004;41(4):381-386. doi:<u>10.1597/02-142.1</u>
- Moher D, Shamseer L, Clarke M, et al. Preferred reporting items for systematic review and meta-analysis protocols (PRISMA-P) 2015 statement. Syst Rev. 2015;4(1):1. doi:10.1186/2046-4053-4-1
- Liberati A, Altman DG, Tetzlaff J, et al. The PRISMA statement for reporting systematic reviews and metaanalyses of studies that evaluate health care interventions: explanation and elaboration. J Clin Epidemiol. 2009;62(10):e1-e34. doi:10.1016/j.jclinepi.2009.06.006
- University of York. Systematic Reviews: CRD's guidance for undertaking reviews in healthcare. York Publishing Services Ltd; 2009. Accessed July 29, 2022. <u>https://www. york.ac.uk/media/crd/Systematic_Reviews.pdf</u>
- 19. Wells GA, Shea B, O'Connell D, et al. Newcastle-Ottawa quality assessment scale control studies. University of Ottawa; 2014:2. Accessed April 11, 2023. <u>https://www. ohri.ca/programs/clinical_epidemiology/nosgen.pdf</u>
- 20. Higgins JPT, Thomas J, Chandler J, et al, eds. Cochrane Handbook for Systematic Reviews of Interventions version 6.3. Cochrane;2022. Accessed July 29, 2022. www.training.cochrane.org/handbook
- Beaty TH, Wang H, Hetmanski JB, et al. A case-control study of nonsyndromic oral clefts in Maryland. Ann Epidemiol. 2001;11(6):434-442. doi:<u>10.1016/s1047-2797(01)00222-8</u>
- 22. Honein MA, Rasmussen SA, Reefhuis J, et al. Maternal smoking and environmental tobacco smoke exposure and the risk of orofacial clefts. Epidemiology. 2007;18(2):226-233. doi:10.1097/01.ede.0000254430.61294.c0
- 23. Jianyan L, Zeqiang G, Yongjuan C, Kaihong D, Bing D, Rongsheng L. Analysis of interactions between genetic variants of BMP4 and environmental factors with nonsyndromic cleft lip with or without cleft palate susceptibility. Int J Oral Maxillofac Surg. 2010;39(1):50-56. doi:10.1016/j.ijom.2009.10.010
- 24. Leite IC, Koifman S. Oral clefts, consanguinity, parental tobacco and alcohol use: a case-control study in Rio de Janeiro, Brazil. Braz Oral Res. 2009;23(1):31-37. doi:10.1590/s1806-83242009000100006

- 25. Li L, Zhu GQ, Meng T, et al. Biological and epidemiological evidence of interaction of infant genotypes at Rs7205289 and maternal passive smoking in cleft palate. Am J Med Genet A. 2011;155A(12):2940-2948. doi:<u>10.1002/</u> <u>ajmg.a.34254</u>
- 26. Li Z, Liu J, Ye R, Zhang L, Zheng X, Ren A. Maternal passive smoking and risk of cleft lip with or without cleft palate. Epidemiology. 2010;21(2):240-242. doi:<u>10.1097/</u> <u>EDE.0b013e3181c9f941</u>
- 27. Lie RT, Wilcox AJ, Taylor J, et al. Maternal smoking and oral clefts: the role of detoxification pathway genes. Epidemiology. 2008;19(4):606-615. doi:10.1097/ EDE.0b013e3181690731
- Mirilas P, Mentessidou A, Kontis E, et al. Parental exposures and risk of nonsyndromic orofacial clefts in offspring: a case-control study in Greece. Int J Pediatr Otorhinolaryngol. 2011;75(5):695-699. doi:<u>10.1016/j.</u> <u>ijporl.2011.02.018</u>
- 29. Wang W, Guan P, Xu W, Zhou B. Risk factors for oral clefts: a population-based case-control study in Shenyang, China. Paediatr Perinat Epidemiol. 2009;23(4):310-320. doi:10.1111/j.1365-3016.2009.01025.x
- 30. Zhang B, Jiao X, Mao L, Xue J. Maternal cigarette smoking and the associated risk of having a child with orofacial clefts in China: a case-control study. J Craniomaxillofac Surg. 2011;39(5):313-318. doi:10.1016/j.jcms.2010.07.005
- 31. Chowchuen B, Surakunprapha P, Winaikosol K, Punyavong P, Kiatchoosakun P, Pradubwong S. Birth prevalence and risk factors associated with CL/P in Thailand. Cleft Palate Craniofac J. 2021;58(5):557-566. doi:10.1177/1055665620956896
- 32. Dien VHA, McKinney CM, Pisek A, Pitiphat W. Maternal exposures and risk of oral clefts in South Vietnam. Birth Defects Res. 2018;110(6):527-537. doi:10.1002/ bdr2.1192
- Goveas SR, Savitha NS. Role of environmental factors in the etiology of non-syndromic cleft lip palate. Int J Sci Study. 2017;4(12):21-26. doi:10.17354/ijss/2017/89
- 34. Hao Y, Tian S, Jiao X, et al. Association of parental environmental exposures and supplementation intake with risk of nonsyndromic orofacial clefts: a case-control study in Heilongjiang Province, China. Nutrients. 2015;7(9):7172-7184. doi:10.3390/nu7095328
- 35. Hoyt AT, Canfield MA, Romitti PA, et al. Associations between maternal periconceptional exposure to secondhand tobacco smoke and major birth defects. Am J Obstet Gynecol. 2016;215(5):613.e1-613.e11. doi:10.1016/j.ajog.2016.07.022
- 36. Junaid M, Narayanan MBA, Jayanthi D, Kumar SGR, Selvamary AL. Association between maternal exposure to tobacco, presence of TGFA gene, and the occurrence of oral clefts. A case control study. Clin Oral Investig. 2018;22(1):217-223. doi:10.1007/s00784-017-2102-6
- 37. Kummet CM, Moreno LM, Wilcox AJ, et al. Passive

smoke exposure as a risk factor for oral clefts-A large international population-based study. Am J Epidemiol. 2016;183(9):834-841. doi:<u>10.1093/aje/kwv279</u>

- 38. McKinney CM, Pisek A, Chowchuen B, et al. Case-control study of nutritional and environmental factors and the risk of oral clefts in Thailand. Birth Defects Res A Clin Mol Teratol. 2016;106(7):624-632. doi:10.1002/bdra.23505
- Regina Altoé S, Borges ÁH, Neves ATSC, et al. Influence of parental exposure to risk factors in the occurrence of oral clefts. J Dent (Shiraz). 2020;21(2):119-126. doi:10.30476/DENTJODS.2019.77620.0
- 40. Chevrier C, Bahuau M, Perret C, et al. Genetic susceptibilities in the association between maternal exposure to tobacco smoke and the risk of nonsyndromic oral cleft. Am J Med Genet A. 2008;146A(18):2396-2406. doi:10.1002/ajmg.a.32505
- Sato Y, Yoshioka E, Saijo Y, et al. Population attributable fractions of modifiable risk factors for nonsyndromic orofacial clefts: a prospective cohort study from the Japan environment and children's study. J Epidemiol. 2021;31(4):272-279. doi:10.2188/jea.JE20190347
- 42. Pi X, Li Z, Jin L, et al. Secondhand smoke during the periconceptional period increases the risk for orofacial clefts in offspring. Paediatr Perinat Epidemiol. 2018;32(5):423-427. doi:10.1111/ppe.12497
- 43. World Health Organization. New brief outlines devastating harms from tobacco use and exposure to second-hand tobacco smoke during pregnancy and throughout childhood Report calls for protective policies. World Health Organization. March 16, 2021. Accessed July 29, 2022. https://www.who.int/news/item/16-03-2021-new-brief-outlines-devastating-harms-from-tobacco-use-and-exposure-to-second-hand-tobacco-smoke-during-pregnancy-and-throughout-childhood
- 44. Peterson LA, Hecht SS. Tobacco, e-cigarettes, and child health. Curr Opin Pediatr. 2017;29(2):225-230. doi:10.1097/MOP.00000000000456
- 45. Chan YL, Saad S, Machaalani R, et al. Maternal cigarette smoke exposure worsens neurological outcomes in adolescent offspring with hypoxic-ischemic injury. Front Mol Neurosci. 2017;10:306. doi:<u>10.3389/ fnmol.2017.00306</u>
- 46. Gupta PC, Subramoney S. Smokeless tobacco use and risk of stillbirth: a cohort study in Mumbai, India. Epidemiology. 2006;17(1):47-51. doi:<u>10.1097/01.</u> <u>ede.0000190545.19168.c4</u>
- Berthiller J, Sasco AJ. Smoking (active or passive) in relation to fertility, medically assisted procreation and pregnancy. Article in French. J Gynecol Obstet Biol Reprod. 2005;34(suppl 1):3s47-3s54. doi:10.1016/ <u>S0368-2315(05)82970-9</u>
- Sawyer DE, Aitken RJ. Male-mediated developmental defects and childhood disease. Reprod Med Rev. 2000;8(2):107-126. doi:10.1017/S0962279900000211
- 49. Zhou Q, Zhang S, Wang Q, et al. Association between

preconception paternal smoking and birth defects in offspring: evidence from the database of the National Free Preconception Health Examination Project in China. BJOG. 2020;127(11):1358-1364. doi:10.1111/1471-0528.16277

- 50. Fell M, Russell C, Medina J, et al. The impact of changing cigarette smoking habits and smoke-free legislation on orofacial cleft incidence in the United Kingdom: evidence from two time-series studies. PLoS One. 2021;16(11):e0259820. doi:10.1371/journal. pone.0259820
- Schiavone S, Anderson C, Mons U, Winkler V. Prevalence of second-hand tobacco smoke in relation to smokefree legislation in the European Union. Prev Med. 2022;154:106868. doi:10.1016/j.ypmed.2021.106868
- Komiyama M, Hasegawa K. Coronavirus disease 2019: psychological stress and cardiovascular diseases. Eur Cardiol. 2021;16:e33. doi:10.15420/ecr.2021.10
- 53. Aljohani MM, Aluqmani AY, Alrehaili EA, Almohammadi MN, Alahmadi HA, Mohamed KG. The effect of public health measures during the COVID-19 pandemic on smoking dependence & passive smoking. J Pharm Res Int. 2021;33(53A):167-180. doi:10.9734/jpri/2021/ v33i53A33649
- 54. Sabbagh HJ, Abdelaziz W, Quritum M, et al. Cigarettes' use and capabilities-opportunities-motivation-forbehavior model: a multi-country survey of adolescents and young adults. Front Public Health. 2022;10:875801. doi:10.3389/fpubh.2022.875801

CONFLICTS OF INTEREST

The authors have completed and submitted the ICMJE Form for Disclosure of Potential Conflicts of Interest and none was reported.

FUNDING

There was no source of funding for this research.

ETHICAL APPROVAL AND INFORMED CONSENT Ethical approval and informed consent were not required for this study.

DATA AVAILABILITY

The data supporting this research are available from the authors on reasonable request.

AUTHORS' CONTRIBUTIONS

Conceptualization, methodology, study design, data analysis, writing and editing of final draft: HJS and KKB. Methodology, study design, article gathering and writing of proposal and first draft: HMJ, ASB, SMA, OAA and KMA. Data analysis and writing and editing of final draft: MHAH.

PROVENANCE AND PEER REVIEW

Not commissioned; externally peer reviewed.